Double robustness of a model-X conditional independence test

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Ziang Niu





Niu*, Chakraborty*, Dukes, & Katsevich. Reconciling model-X and doubly robust approaches to conditional independence testing. *Annals of Statistics*, 2024.





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Oliver Dukes



Eugene Katsevich

Statistical task: Test whether a response variable $\mathbf{Y} \in \mathbb{R}$ is associated with a predictor variable $\mathbf{X} \in \mathbb{R}$ when controlling for covariates $\mathbf{Z} \in \mathbb{R}^p$, given *n* i.i.d. samples (X_i, Y_i, Z_i) from a joint distribution $\mathscr{L}_n(\mathbf{X}, \mathbf{Y}, \mathbf{Z})$.

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Hypothesis formulation: In the joint distribution $\mathscr{L}_n(\mathbf{X}, \mathbf{Y}, \mathbf{Z})$, test the null hypothesis of conditional independence (CI):

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- **Hypothesis formulation:** In the joint distribution $\mathscr{L}_n(\mathbf{X}, \mathbf{Y}, \mathbf{Z})$, test the null
 - $H_0^{\mathsf{CI}}: \mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z}.$

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 - This turns out to be a very challenging problem!

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A test with Type-I error control must protect against too many sneaky ways Z can affect both X and Y.

Given a set of regularity conditions \mathscr{R}_n on \mathscr{L}_n , one can only hope to control Type-I error over the smaller null hypothesis

$$H_0: H_0^{\mathsf{Cl}} \cap \mathscr{R}_n$$



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 n^{\bullet}

The model-X (MX) assumption (Candès et al '18)

Assume we know the conditional distribution $\mathscr{L}_n(\mathbf{X} \mid \mathbf{Z})$ exactly, i.e. $\mathscr{R}_n \equiv \{\mathscr{L}_n : \mathscr{L}_n(\mathbf{X} \mid \mathbf{Z}) = \mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})\},\$ where $\mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})$ is the given conditional distribution.

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Powerful CI tests available under the MX assumption: conditional randomization test (CRT) for single testing and MX knockoffs for multiple testing.²

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4. Reject if $T_n > C_n$.



The conditional randomization test (CRT) **Conditional randomization test Properties** 1. Compute test stat $T_n \equiv T_n(X, Y, Z)$; 2. For b = 1, ..., B, • Draw $\tilde{X}_{i}^{(b)} \stackrel{\text{ind}}{\sim} \mathscr{L}_{n}^{*}(X_{i} \mid \mathbb{Z} = Z_{i});$ • Recompute $\tilde{T}_{n}^{(b)} \equiv T_{n}(\tilde{X}^{(b)}, Y, Z);$ $C_{n} \equiv \mathbb{Q}_{1-\alpha}[\{T_{n}, \tilde{T}_{n}^{(1)}, \dots, \tilde{T}_{n}^{(B)}\}];$

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Remark:

Test statistic choice impacts power;¹ often employs penalized regression or black-box machine learning.



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 - 1. Either out of sample, based on extra unlabeled pairs (X_i, Z_i) ,
 - 2. Or in sample, based on the same data used for testing (more common).

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If $\mathscr{L}_n(\mathbf{X} \mid \mathbf{Z})$ obtained from well-specified OLS based on N auxiliary samples, then $\mathbb{P}[\text{false rejection}] \leq \alpha + O_p\left(\sqrt{\frac{n \cdot \dim(\mathbf{Z})}{N}}\right)$



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CRT properties if $N \gg n \cdot \dim(\mathbb{Z})$ Asymptotic

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Case 2: $\mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})$ learned in sample



MX method applied as if $\mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})$ were known (more common in practice).

- **Theory:** For worst-case test statistics, no hope for Type-I error control.¹ Beyond that, few existing results.
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How robust are MX methods when $\mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})$ learned in sample?



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Open question:

We study this question in the context of a specific MX method: the dCRT.

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The distilled CRT and its in-sample approximation Let $\mu_{n,x}(\mathbf{Z}) \equiv \mathbb{E}_{\mathscr{L}_n}[\mathbf{X} \mid \mathbf{Z}]$ and $\mu_{n,y}(\mathbf{Z}) \equiv \mathbb{E}_{\mathscr{L}_n}[\mathbf{Y} \mid \mathbf{Z}].$

Let $\mu_{n,x}(\mathbf{Z}) \equiv \mathbb{E}_{\mathscr{L}_n}[\mathbf{X} \mid \mathbf{Z}]$ and $\mu_{n,y}(\mathbf{Z})$

The dCRT¹ is an instance of the CRT, with $T_n(X, Y, Z) \equiv \frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i)$

is known by the MX assumption.

$$) \equiv \mathbb{E}_{\mathscr{L}_{n}}[\mathbf{Y} \mid \mathbf{Z}].$$

$$(X_i - \mu_{n,x}(Z_i))(Y_i - \hat{\mu}_{n,y}(Z_i)),$$

where $\hat{\mu}_{n,v}(\mathbf{Z})$ is obtained via in-sample machine learning of Y on Z and $\mu_{n,x}(\mathbf{Z})$

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The dCRT¹ is an instance of the CRT, with $T_n(X, Y, Z) \equiv \frac{1}{\sqrt{n}} \sum_{i=1}^n (A_i)^{(n)}$

is known by the MX assumption.

The approximate dCRT with $\widehat{\mathscr{D}}_{n}(X \mid Z)$ learned in sample is the same, except $T_n(X, Y, Z) \equiv \frac{1}{r} \sum_{n=1}^{n} \sum_{n=$ $n_{i=1}$

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The approximate dCRT with $\widehat{\mathscr{D}}_{n}(X \mid Z)$ learned in sample is the same, except $T_n(X, Y, Z) \equiv \frac{1}{7} \sum_{n=1}^{n} \sum_{n=$ $\sqrt{n} \sum_{i=1}^{n}$

$$) \equiv \mathbb{E}_{\mathscr{L}_{n}}[\mathbf{Y} \mid \mathbf{Z}].$$

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$$\begin{array}{l} (X_i - \hat{\mu}_{n,x}(Z_i))(Y_i - \hat{\mu}_{n,y}(Z_i)). \\ \text{Resampling distribution} \\ \text{changed to } \hat{\mathscr{L}}(X_i \mid Z_i) \end{array}^{1\text{Liu et al '22}} \end{array}$$



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where $\hat{\mu}_{n,y}(\mathbf{Z})$ is obtained via in-sample machine learning of Y on Z and $\mu_{n,x}(\mathbf{Z})$ is known by the MX assumption.



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Resampling distribution changed to $\widehat{\mathscr{L}}(X_i | Z_i)$ ¹Liu et al '22 10



Is dCRT robust to in-sample learning?

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Estimate of E[YIZ] — Intercept–only

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Case 2: $\mathbb{E}[Y \mid Z]$ estimated decently: $\widehat{\mu}_{n.v}(\mathbf{Z})$ obtained via lasso of *Y* on *Z*.



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No hope for good inference with poor estimate for $\mathbb{E}[Y \mid Z]$.



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- We found that for dCRT, better estimation of $\mathscr{L}(\mathbf{Y} \mid \mathbf{Z})$ compensates for errors in the estimation of $\mathscr{L}(X \mid Z)$. This is a double robustness phenomenon!
- We claim that the dCRT itself is doubly robust!
- In fact, we claim that the dCRT is asymptotically equivalent to the doubly robust generalized covariance measure (GCM) test (Shah and Peters, 2020).



1. Fit an approximation $\hat{\mu}_{n,x}(\mathbf{Z})$ of $\mu_{n,x}(\mathbf{Z}) \equiv \mathbb{E}_{\mathscr{L}_n}[\mathbf{X} \mid \mathbf{Z}]$ via machine learning;

- 3. Compute $T_n(X, Y, Z) \equiv \frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i \hat{\mu}_{n,x}(Z_i))(Y_i \hat{\mu}_{n,y}(Z_i))$

- 3. Compute $T_n(X, Y, Z) \equiv \frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i \hat{\mu}_{n,x}(Z_i))(Y_i \hat{\mu}_{n,y}(Z_i))$ 4. Compute $(S_n^{\text{GCM}})^2(X, Y, Z) \equiv \mathbb{V}\{(X_i \hat{\mu}_{n,x}(Z_i))(Y_i \hat{\mu}_{n,y}(Z_i))\}$

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- 5. Reject if $\frac{T_n(X, Y, Z)}{S_n^{\text{GCM}}(X, Y, Z)} > z_{1-\alpha}.$



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- 5. Reject if $\frac{T_n(X, Y, Z)}{S_n^{GCM}(X, Y, Z)} > z_{1-\alpha}$. Asymptotic threshold, rather than resampling-based.



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If $\text{RMSE}(\hat{\mu}_{n,x}) = o_P(1)$, $\text{RMSE}(\hat{\mu}_{n,y}) = o_P(1)$, $\text{RMSE}(\hat{\mu}_{n,x}) \cdot \text{RMSE}(\hat{\mu}_{n,y}) = o_P(n^{-1/2})$ for each $\mathscr{L}_n \in H_0 \equiv H_0^{\mathsf{CI}} \cap \mathscr{R}_n$, then GCM test has asymptotic Type-I error control.
Double robustness of the GCM test

The GCM test is doubly robust (Shah and Peters '20)

e.g. $s = o(\sqrt{n}/\log p)$.

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These rates allow for high-dimensional regressions in the "consistency regime,"

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10 samples



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25 samples -2.5 0.0 2.5 5.0 **Resampled statistic**

250 samples







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We proved a new conditional CLT for triangular arrays!









Theorem (Niu et al '24; informal). Assume

- 1. $\text{RMSE}(\hat{\mu}_{n,x}) = o_P(1), \text{RMSE}(\hat{\mu}_{n,y})$
 - $\text{RMSE}(\hat{\mu}_{n,x}) \cdot \text{RMSE}(\hat{\mu}_{n,y}) = o_P(n^{-1/2}).$
- 2. The estimated variances are "consistent" in some sense.

$$p_{P}(1) = o_{P}(1),$$

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Then, for any $\mathscr{L}_n \in H_0$, the dCRT is asymptotically equivalent to the GCM test.



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Fitting $\mathbb{E}[\mathbf{Y} \mid \mathbf{Z}]$ improves not just power; it improves robustness as well.





(Previous work)



CRT properties under **MX** assumption ($\mathscr{L}_n^*(\mathbf{X} \mid \mathbf{Z})$ known)

Finite-sample Type-I error control

- No assumptions on $\mathscr{L}_n(\mathbf{Y} \mid \mathbf{Z})$ \bullet
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dCRT requires a large number of resamples to obtain accurate small p-values.

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We extend saddlepoint approximation theory to approximate the conditional tail



spaCRT is completely resampling-free and almost as fast as GCM!



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Discussion
Take-home message:

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• When $\mathscr{L}_n(\mathbf{X} \mid \mathbf{Z})$ fit in sample, MX inference is like doubly robust inference; Type-I error control possible, but both $\widehat{\mathscr{L}}_n(\mathbf{X} \mid \mathbf{Z})$ and $\widehat{\mathscr{L}}_n(\mathbf{Y} \mid \mathbf{Z})$ matter.

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- Moving beyond the "consistency regime," e.g. to proportional asymptotics

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Thank you! Questions?



Simple numerical simulation:

- $\mathscr{L}(\mathbf{Z}) \sim N(0,I); \ \mathscr{L}(\mathbf{X} \mid \mathbf{Z}) = N(\mathbf{Z}^T \beta, 1); \ \mathscr{L}(\mathbf{Y} \mid \mathbf{Z}) = N(\mathbf{Z}^T \beta, 1);$
- n = 1600, p = 400, β has 5 nonzero elements;
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improves robustness of dCRT.

Corollary (Niu et al '24; informal). Assume

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under Shah and Peters's conditions, so GCM is also most powerful.

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The GCM statistic $T_n(X, Y, Z)$ is asymptotically equivalent to $T_n^{\text{oracle}}(X, Y, Z)$

Testing the CI null $H_0^{CI} \cap \mathscr{R}_n$ is not the same as testing the semiparametric null

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Therefore, any test controlling Type-I error on $H_0^{\mathsf{CI}} \cap \mathscr{R}_n$ must also control Type-I error on H_0^{SP} , and so its power is bounded above by that of the GCM test.



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Theorem (Niu et al '24; informal). Under Shah and Peters's conditions, the GCM test of $H_0^{\mathsf{CI}} \cap \mathscr{R}_n$ is asymptotically most powerful against $H_1^{\mathsf{SP}} : \mathbf{Y} = \mathbf{X}h/\sqrt{n} + g(\mathbf{Z}) + \epsilon$.




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For alternatives with interactions or heteroskedasticity:

• We would not expect GCM (or dCRT) to be optimal, and alternative methods may have better power.¹

¹Scheidegger et al '21, Zhong et al '21, Lundborg et al. '22

Numerical simulations: Design

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Data-generating model:

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Data-generating model:

$\mathscr{L}(\mathbf{Z}) = N(0, \Sigma(\rho)), \ \mathscr{L}(\mathbf{X} \mid \mathbf{Z}) = N(\mathbf{Z}^T \beta, 1), \ \mathscr{L}(\mathbf{Y} \mid \mathbf{X}, \mathbf{Z}) = N(\mathbf{X}\theta + \mathbf{Z}^T \beta, 1),$





Parameters ν and θ control degree of confounding and signal strength.



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Methods compared:



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dCRT and GCM (with lasso and post-lasso)



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Methods compared:

- dCRT and GCM (with lasso and post-lasso)
- Maxway CRT (implemented with data splitting)



n = 200; p = 400; ρ = 0.4

s = 20 s = 5

Marginal association between X and Y (v)

- dCRT (LASSO) GCM (LASSO) Maxway CRT
- dCRT (PLASSO) - GCM (PLASSO)

s = 80

n = 200; p = 400; ρ = 0.4



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Takeaways

GCM and dCRT perform similarly, \bullet consistent with asymptotic theory.

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- Lasso-based methods can have very \bullet inflated Type-I error in difficult settings.

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Remark

We expect, for smaller samples sizes or lacksquaremore discrete data, that dCRT can have better Type-I error control than GCM.

n = 200; p = 400; ρ = 0.4



dCRT (PLASSO) - - GCM (PLASSO) - -



= 80

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Note: All methods subjected to "oracle calibration" for fair power comparison.

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GCM (LASSO) dCRT (LASSO) ____

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Maxway CRT

30

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Takeaways

Maxway CRT

n = 200; p = 400; ρ = 0.4



- dCRT (LASSO) - GCM (LASSO) - Maxv

- - dCRT (PLASSO) - - GCM (PLASSO)

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Takeaways

• GCM tends to outperform dCRT.

Maxway CRT

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Takeaways

- GCM tends to outperform dCRT.
- Lasso outperforms post-lasso, suggesting bias-variance trade-off.

Maxway CRT

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- dCRT (LASSO) - GCM (LASSO) - Maxv

- - dCRT (PLASSO) - - GCM (PLASSO)

Maxway CRT

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- GCM tends to outperform dCRT.
- Lasso outperforms post-lasso, suggesting bias-variance trade-off.
- Maxway CRT has lowest power, due to data splitting. Better performance in separate semi-supervised simulation.